

Impact of COVID-19 Activity Restrictions on Air Pollution: Methodological Considerations in the Economic Valuation of the Long-Term Effects on Mortality

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Abstract – This article offers an approach incorporating latency into the process for evaluating long-term mortality and into its economic valuation, following a temporary impact. It is applied to the effects of COVID-19 activity restrictions, in the spring of 2020, on ambient air pollution in France. These effects are evaluated in terms of Life Years Gained (LYG) and in monetary terms for two air pollution indicators. This approach is compared to a standard estimate on the basis of difference. It gives results that are lower by a factor of 3.7 to 5.5 for LYG and, on account of the additional effect of discounting, gives an economic valuation that is lower by a factor of 4.7 to 6.9. These results show that an adapted valuation of the long-term health benefits, then their translation into monetary terms, is essential in order to compare the long-term consequences of temporary exogenous impacts or policies.

JEL classification: C18, I1, Q51, Q53

Keywords: COVID-19, long-term mortality, activity restrictions, air pollution, economic valuation

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The author wishes to thank Lucie Adélaïde, Guillaume Boulanger, Augustin Colette, Sylvia Medina, Mathilde Pascal and Vèrène Wagner for their very useful contributions, and two anonymous reviewers for their comments and suggestions. This work was supported by the French National Research Agency Grant ANR-17-EURE-0020, and by the Excellence Initiative of Aix-Marseille University - A*MIDEX.

Received in October 2021, accepted in March 2022. Translated from « Impact sur la pollution de l'air des restrictions d'activité liées à la Covid-19 : apports méthodologiques dans l'évaluation économique des effets de long terme sur la mortalité »

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Citation: Chanel, O. (2022). Impact of COVID-19 Activity Restrictions on Air Pollution: Methodological Considerations in the Economic Valuation of the Long-Term Effects on Mortality. *Economie et Statistique / Economics and Statistics*, 534-35, 103–118 (First published online: May 2022).

doi: 10.24187/ecostat.2022.534.2081

Beyond the direct impacts on morbidity and mortality, COVID-19 led to radical changes of lifestyle for the population as of March 2020. As in most countries (Liu *et al.*, 2021), France imposed activity restrictions in spring 2020. The many negative consequences of these – social, educational, professional and health – are likely to increase the socio-economic inequalities within the population and cannot yet be fully evaluated (Bambra *et al.*, 2020; Tisdell, 2020; Brodeur *et al.*, 2021b). From a health standpoint, the impacts include mental health damage, a decrease in physical activity, loss of opportunities in medical terms linked to the inability to monitor chronic illness and surgery cancellations, changes to eating habits, increased exposure to indoor air pollution or reduced well-being linked to lockdown (Brodeur *et al.*, 2021a; Hrynich *et al.*, 2021; Le & Nguyen, 2021; Molina-Montes *et al.*, 2021). Some consequences of lockdown were positive, however, as the activity restrictions were accompanied by a drop in the number of road traffic accidents (in France, about 720 fewer deaths and 14,900 fewer injuries in 2020 than in 2019, cf. ONISR – *Observatoire national interministériel à la sécurité routière*, 2021), and reductions in ambient concentrations of certain atmospheric pollutants and the associated health effects.

This article studies the consequences of this reduction on long-term mortality. Short-term mortality has been widely studied (Bherwani *et al.*, 2020; Chen *et al.*, 2020; Wang *et al.*, 2020; Liu *et al.*, 2021; Sannigrahi *et al.*, 2021; Venter *et al.*, 2021), but its long-term counterpart less so (however, see Giani *et al.*, 2020; Adélaïde *et al.*, 2021b; Hao *et al.*, 2021). When it was studied, the results obtained were not adapted to economic valuation. In fact, these studies evaluated the effects on mortality based on two situations – with and without lockdown – and calculated the consequences by considering the difference between these two situations over a given period, *ceteris paribus*. This standard approach based on difference – clear, simple and instructive – is perfectly suited to short-term effects. However, its value is limited for long-term effects as it does not take account of the cumulative nature of the exposure which dictates the distribution of health benefits over time. Therefore, disregarding this latency when evaluating health effects has repercussions on the economic valuation of future benefits, which are amplified by discounting.

As such, we offer an approach that incorporates latency when assessing the effects of a temporary impact on long-term mortality and its economic

valuation. We apply this approach to the drops in concentration of two atmospheric pollutants observed in mainland France in 2020: fine particles PM_{2.5} (aerodynamic diameter less than 2.5 µm) and nitrogen dioxide (NO₂). We find that the standard approach based on difference gives results in terms of Life Years Gained (LYG) that are considerably higher than those obtained using the approach that we offer, by a factor of 3.7 for PM_{2.5} and 5.5 for NO₂. Under the effect of discounting when performing economic valuation, these factors rise to 4.7 and 6.9 respectively. Generally speaking, an adapted valuation of long-term health benefits, then its translation into monetary terms, is essential to allow the economist to compare the long-term consequences of temporary public policies or exogenous impacts such as COVID-19.

We set out the methodology used to evaluate the health and economic impacts, in particular the use of uncertainty (Section 1). We apply this to the impact on long-term mortality of drops in pollution levels resulting from the restrictions related to COVID-19 in spring 2020 (Section 2). The results are shown in Section 3.

1. Methodology for the Economic Valuation of Health Impacts

1.1. Standard Approach Based on Difference

The association between pollution indicators and health indicators is based on statistical models that estimate exposure-response functions. For most pollutants and long-term mortality, these functions are considered linear and non-threshold (WHO, 2021). Therefore, the relative risks (RR) used quantify the variations in mortality in a population when its exposure varies, regardless of the initial exposure level. They are used as a basis for calculating three indicators: the number of premature deaths, the total number of years of life and life expectancy at a given age. The latter two require the use of dynamic mortality tables for the population concerned: the RR of mortality associated with exposure to the pollutant affects the probability of death from any cause, and the synthetic cohort is monitored until its extinction.

Epidemiological studies generally apply a difference-based approach to determine the health effects of a variation in exposure. The RR is then applied to the exposure differential and to the average annual number of deaths, or used to evaluate a number of LYG based on the difference between the evolution of cohorts exposed or not to this exposure variation (for example,

Corso *et al.*, 2019, pp. 46–50 for the methodology). When the variation is permanent, these RR are used to determine the annual long-term impact; when it is temporary, as with lockdown, they determine the total long-term impact. In both cases, the health effects are considered to be immediate.

1.2. Impacts of Latency on the Distribution over Time of Health Gains Following a Temporary Shock

The standard approach based on difference is not, however, adapted to a long-term mortality RR, translating the impact on state of health of a cumulative process, which is not immediate in either its degeneration or its improvement (Leksell & Rabl, 2001; Miller & Hurley, 2003; Rösli *et al.*, 2005; Burnett *et al.*, 2018). We are, therefore, seeking a framework adapted to a drop in exposure which is temporary, and where the long-term health effects would not be immediate.

1.2.1. Literature Review

Epidemiological literature on the effects of air pollution rarely studies this process on account of a lack of data on the change over time of the long-term RR following an exposure modification. Walton (2010) produces a very comprehensive analysis based on three sources: time-based trends taken from epidemiological studies, the biological processes underlying the different types of associated mortality (cardiopulmonary, cardiovascular, respiratory and lung cancer), and certain similar risk factors which are better quantified, such as stopping smoking. Despite the existence of uncertainties, the first two sources confirm a non-immediate effect which stretches over several years on account of the mechanics of deterioration and recovery associated with the health effects, without being able to precisely determine the distribution over time.

This latter may, however, be inferred from data on smoking cessation, an area in which Walton (2010) compiles 22 studies published between 1976 and 2008, which indicate that the mortality of ex-smokers is similar to that of individuals who have never smoked, after a period of abstinence of 10 to 20 years. It is strongly demonstrated that cardiovascular mortality decreases rapidly over the first five years, while maintaining a component that diminishes more gradually up to 20 or 30 years after stopping, whereas lung cancer mortality decreases more gradually over 30 years.

On these bases, and given that the exposure route (inhalation) and target organs (pulmonary

system) are common to tobacco and pollution exposure, several structures for latency distribution have been proposed. Some of these cover a relatively short timeframe: 85% the first year with the remaining 15% over the next six years (Laden *et al.*, 2006), or 25% per year over the first four years (Puett *et al.*, 2009). Other approaches consider a longer time period: uniform distribution over the first 15 years (Krewski *et al.*, 2009), 40% in the first five years, and the remaining 60% over the following 30 years (Walton, 2010); or a decreasing exponential structure with 50% in the first six years and the remainder over the next 40 years (Rösli *et al.*, 2005).

Empirically, analyses of the benefits carried out by the Environmental Protection Agency (US EPA, 2021) have, since 2006, applied a 20-year lag structure: 30% of premature deaths arising during the year following the reduction (the contribution of short-term exposure), 50% spread equally over years 2 to 5 following the reduction (deaths of cardiopulmonary origin) and 20% distributed equally over years 6 to 20 following the reduction (deaths due to pulmonary disease and lung cancer).

Ultimately, we conclude, along with Rabl (2006), that the data available support the impact of atmospheric pollution on mortality in proportion to the integration over time of past concentrations, weighted by a decreasing exponential profile.

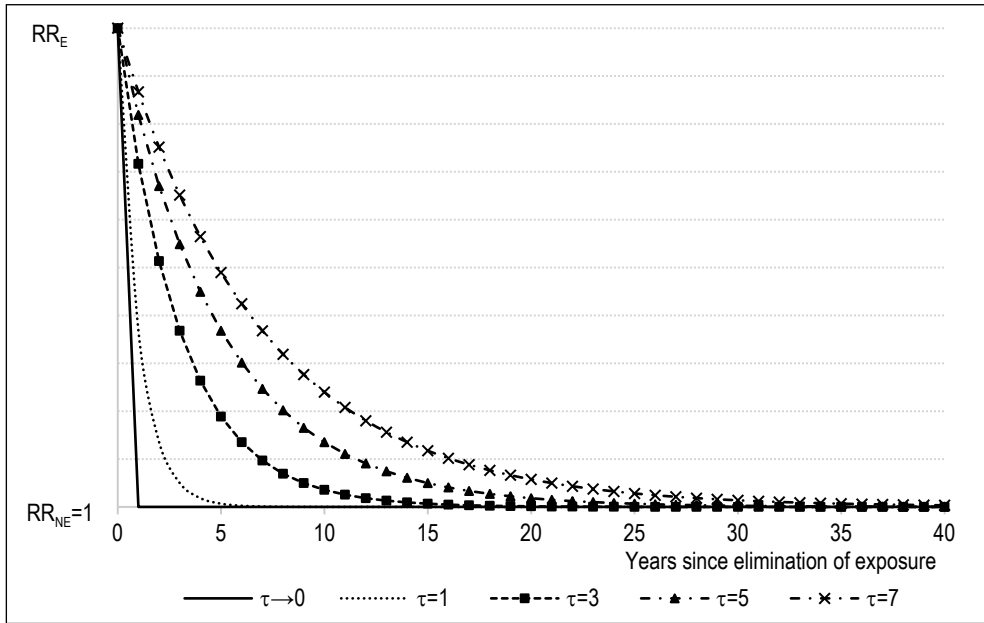
1.2.2. Consideration of Latency for a Permanent Elimination of Exposure

Lightwood & Glantz (1997) thus estimate a negative exponential mortality risk function (like Rösli *et al.*, 2005), based on the meta-analysis of seven studies on the impacts of smoking cessation, which represents an immediate and complete elimination of the risk:

$$RR(t) = RR_{NE} + (RR_E - RR_{NE})e^{\left(-\frac{t}{\tau}\right)} \quad (1)$$

where RR_E is the RR linked to exposure to a risk factor (active smoking in smokers), RR_{NE} the RR associated with no exposure to this factor (absence of smoking in non-smokers), $e(\cdot)$ the exponential function, t the time elapsed since elimination of the exposure (stopping smoking) and τ a parameter > 0 . If $\tau \rightarrow 0$, the impact on the RR is obtained immediately, and concurs with the standard approach based on difference. When τ increases, the time necessary for $RR(t)$ to reach RR_{NE} increases. Figure I represents the change in $RR(t)$ for different values of τ : immediate decrease when τ is close to 0 (solid line); decrease over approximately six years for $\tau = 1$;

Figure I – Change in the relative risk (RR) of mortality following a permanent elimination of exposure, function of τ



over 20 years for $\tau=3$; over 30 years for $\tau=5$; and over 40 years for $\tau=7$.

Some studies on the long-term effects of exposure to atmospheric pollution have adopted and applied this formula (Leksell & Rabl, 2001; Chanel *et al.*, 2006; Rabl, 2006) or its counterpart for air pollution (Röösli *et al.*, 2005; Tainio *et al.*, 2007), favouring epidemiological data specific to the diseases leading to death. They performed a sensitivity analysis on the value of τ , liable to represent the gradual decrease in mortality over the longer term, in order to take account of the associated uncertainties.

1.2.3. Consideration of a Temporary Elimination of Exposure

However, the reduction in exposure is deemed permanent in the case of smoking cessation, whereas we are looking at – to use the expression of Johannesson *et al.* (1997) – the impact of a *blip* on mortality, i.e. a low, immediate and temporary reduction, with a return to the previous exposure level. We are, therefore, adapting the mortality risk function from equation (1) to model this return to the level of RR_E when exposure to the factor is re-established at its initial level (as $t=t_0$). This then gives us, with the previous notations:

$$RR(t) = RR_E + (RR(t_0) - RR_E) e^{\left[\frac{-(t-t_0)}{\tau} \right]} \quad (2)$$

for $t \geq t_0$

Figure II shows the change in $RR(t)$ for a temporary elimination of exposure over five years

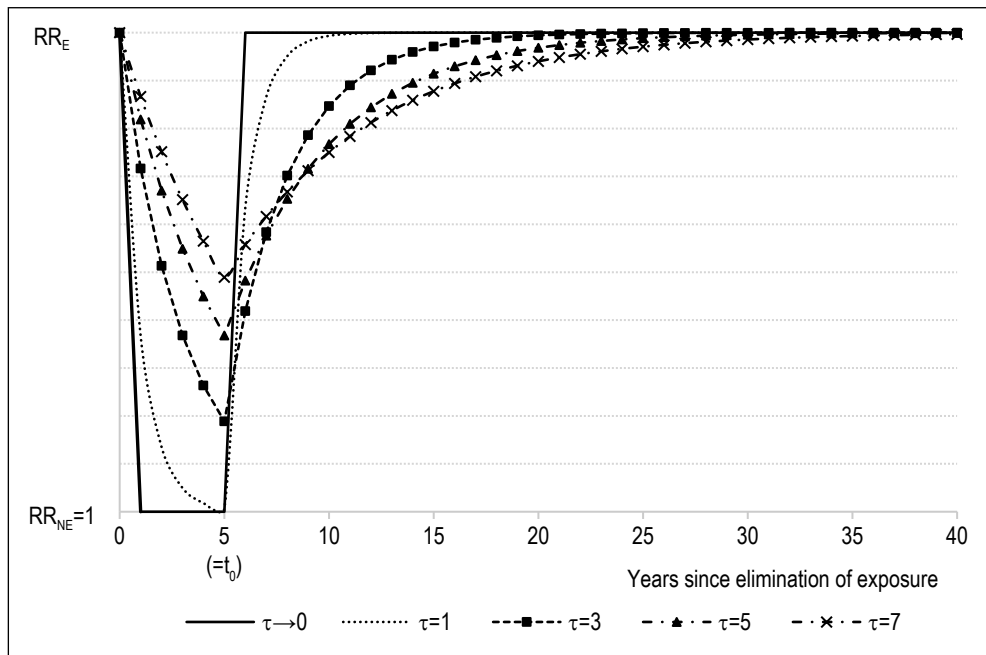
($t_0=5$) and for different values of τ . It shows that the higher the value of τ , the quicker $RR(t)$ drops, to achieve a value at the end of the period during which exposure is eliminated that is closer to RR_{NE} , but that more time is needed to return to the level of RR_E (five years for $\tau=1$ but 35 years for $\tau=7$).

1.2.4. Choice of Value of the Parameter τ

Estimates of τ differ in literature depending on the disease causing the death. With regard to smoking cessation, Lightwood & Glantz (1997) suggest 1.4 for a stroke and 1.6 for an acute myocardial infarction, Leksell (2000) between 4.3 and 6.5 for lung cancer, and Doll *et al.* (1994) between 10 and 15 for a total excess risk of mortality. Leksell & Rabl (2001) find that a good approximation for mortality across all causes is a weighted average where $\tau=1.5$ (weight of 0.3) and $\tau=13$ (weight of 0.7).

With regard to exposure to air pollution, Röösli *et al.* (2005) estimate τ for two interventional studies and obtain 1.1 (for elimination of exposure to the emissions of a steel mill for 13 months) and 9 (for permanent elimination of exposure to coal, but a follow-up of only six years). For their own study, they choose a central value $\tau=5$ with a sensitivity analysis ranging from $\tau \rightarrow 0$ to $\tau=10$.

Ultimately, we have chosen a central value of $\tau=3$, which corresponds approximately to the empiric distribution used by the US EPA (2021). Indeed, Figure I indicates that 30% of the risk variation ($RR_E - RR_{NE}$) is obtained in the first

Figure II – Change in the relative risk (RR) of mortality following a temporary elimination of exposure, function of τ 

year, 50% for the period from 2 to 5 years and 20% for the period from 6 to 20 years. We have chosen the values $\tau=1$ and $\tau=5$ as the uncertainty interval.

1.3. Impacts of Latency on the Economic Valuation of Mortality

From an economic point of view, incorporating latency and distribution over time in LYG involves the use of discounting to express future monetary flows as current values, whether through years of life (Hammit, 2007; Jones-Lee *et al.*, 2015) or the valuation of future monetary gains (US EPA, 2021). Thus, using the LYG distribution over time, we obtain the following total economic valuation:

$$\text{Total economic valuation} = \sum_{t=1}^{120} LYG_t \text{VOLY} (1+\delta)^{-t} \quad (3)$$

where LYG_t represents the number of LYG on date t , VOLY the value of a life year, and δ the discount rate, the latter two having to be chosen. The upper limit of the sum is set at 120 years, the maximum age that guarantees extinction of the cohort.

1.4. Accounting for Uncertainties

The economic valuation of the effects of exposure of the population to ambient air must take into account the accumulated uncertainties that mainly arise from three sources.

Firstly, the uncertainties in the characterisation of population exposure, mainly due to

the measurement of concentrations and of the exposure (observed), and to the modelling of the counterfactual exposure (not observed). The quality of the modelling depends on the quality of the input data (emissions inventories, land use data, geographical distribution of the population, meteorological data, etc.), the topography of the area studied, the availability of measurement data, etc., making the uncertainty spatially heterogeneous.

Next, epidemiological uncertainties concern the quality of the health data, the choice of a risk-exposure function (functional form, thresholds) or an RR, and their transposability to the population studied, which depends on way of life, climate or the nature of the emission sources. Part of this uncertainty is provided by the confidence interval, generally 95% (95% CI) around the central RR value. This latter is derived from econometric regressions on data pairs representing the exposure levels and health effects observed, such that the associated uncertainty reflects the statistical variability specific to the relationship between exposure and health effect. We note that as the RR are more frequently calculated based on urban rather than rural populations, the uncertainty is likely to be higher for the latter. Although the value of τ that we select in equations (1) and (2) is based on our analysis of epidemiological knowledge and practice, and not on an objective statistical estimate, this choice does convey an underlying epidemiological uncertainty.

Finally, the quantification of economic uncertainties differs as the underlying knowledge is more subjective than scientific, leading to an approach that is more normative than positive. It is based on the unit monetary values used and technical parameters such as the discount rate. These uncertainties are generally accounted for through a triangular probability distribution (Chanel *et al.*, 2014; Rabl *et al.*, 2014), and/or the construction of a range from an empirical standard deviation under an assumption of normality. For example, CAFE (2005) proposes $\pm 33\%$, which corresponds to a variation of approximately one standard deviation around the mean for normal distribution.

These three types of uncertainty are generally considered either independently or jointly by integrating their respective sources in a Monte-Carlo simulation approach, preferable from a methodological standpoint. A more complex analysis can also be performed by breaking down each source and assigning it a specific distribution (Rabl *et al.*, 2014).

2. Application to the Activity Restrictions Related to COVID-19

A quantitative health impact assessment (HIA) conducted by *Santé publique France* has estimated the impact on long-term mortality of the reductions in levels of PM_{2.5} and NO₂ observed in mainland France during lockdown (Adélaïde *et al.*, 2021b; Medina *et al.*, 2021). We present this methodology briefly, along with our own approach (2.1), before addressing the elements necessary for the economic valuation (2.2) and then for accounting for uncertainties (2.3).

2.1. Evaluation of Health Effects

2.1.1. Modelling of Population Exposure

The first step estimates the difference between the actual exposure of the population to the pollution indicators PM_{2.5} and NO₂ during the periods of strict lockdown (from 16 March to 11 May 2020) and the gradual relaxation of measures (from 11 May to 22 June 2020), and that observed in the absence of these lockdown measures. The latter models the air quality using the CHIMERE chemistry-transport model (co-developed by Ineris and CNRS) on the basis of European scenarios adapted for France by CITEPA (*Centre interprofessionnel technique d'études de la pollution atmosphérique*). The air pollution data are taken from the French approved air quality measurement network. The methodology used is similar to that mobilised for

the Ineris air quality map library.¹ Using population data from the 35,228 communes of mainland France (according to the 2018 communes list), exposure is calculated per grid measuring approximately 4 km by 4 km. The concentration values of the different model grids present within the territory of a commune are then weighted according to the population size defined for each grid. Ultimately, this allows us to calculate the average exposure observed during lockdown, weighted at communal level, and to model that which would have been observed in the absence of any lockdown measures. Calculated as an annual average over the period from 1 July 2019 to 30 June 2020, this represents a drop of 2.9% for PM_{2.5} and 4.7% for NO₂.

2.1.2. Estimate of the health effects for a 10 µg.m⁻³ increase

Medina *et al.* (2021) propose two long-term RR for all causes mortality applying to the population aged 30 years and over. For PM_{2.5}, the RR is 1.15 (95% CI: 1.05-1.25) based on 22 European cohorts from the ESCAPE project and one French cohort (Pascal *et al.*, 2016). It is slightly above the values found by Pope *et al.* (2020): 1.09 (1.07-1.11) taken from 75 international studies, and 1.12 (1.06-1.19) obtained from 10 European studies. The difference may emanate from the exposure method and/or the particular composition, and we favour the RR defined by Medina *et al.* (2021). For NO₂, the long-term mortality RR adopted is 1.023 (1.008-1.037), based on 11 Western studies (PHE, 2018), which is also the value selected by the WHO in its latest guidelines (WHO, 2021). It is comparable to the meta-analyses of Huangfu & Atkinson (2020), with 1.02 (1.01-1.04) over 24 studies, or of Stieb *et al.* (2021), with 1.025 (1.012-1.038) across 53 international studies.

2.1.3. Standard Approach Based on Difference and Approach Taking Latency into Consideration

The two approaches study the impact on mortality of a further reduction of the average exposure of the population to PM_{2.5} and NO₂ over the period from 1 July 2019 to 30 June 2020, following the lockdown measures.

The standard approach based on difference, mobilised in Medina *et al.* (2021), applies the RR to the exposure differential calculated during this period. It thus calculates the number of LYG

1. <https://www.ineris.fr/fr/recherche-appui/risques-chroniques/mesure-prevision-qualite-air/20-ans-evolution-qualite-air>

based on the difference between the development of two fictitious cohorts, one exposed to this exposure variation from the age of 30 years, and the other not exposed.

The approach that we propose applies a reduced risk of mortality obtained from equation (1) for this period to the exposed cohort, followed by a return to the previous exposure level in accordance with equation (2). We are, therefore, modifying the conditional probability of death in the population aged 30 years and over, and following the cohorts to extinction, recording the LYG on each date. In practice, we are using the most recent mortality tables by gender (INSEE, 2018), expressed for two fictitious cohorts of 100,000 births. We then approximate LYG_t , the number of LYG in the French population on each date t , replacing the 100,000 initial births in these fictitious cohorts with the actual numbers of births by gender in France (349,105 female and 364,924 male in 2019, the last known year).

2.2. Evaluation of Economic Effects

The monetary valuation of mortality – always delicate – relies on a standard framework adopted in New-Ext (2004), CAFE (2005), Aphekom (2011) and by the European Environment Agency (Schucht *et al.*, 2021). It is based on the choice of a Value of Prevented Fatality (VPF²) and a Value of a Life Year (VOLY),³ employing three main methods (box). As our analysis is based on variations in number

of LYG, it therefore requires the adoption of a VOLY. The latter can be obtained by derivation from a VPF considered as a flow of VOLY discounted over the remaining life expectancy (Viscusi *et al.*, 1997; Leksell & Rabl, 2001) or by direct estimation in a contextual study of stated preferences. A discount rate δ must also be adopted to value the (future) flows of LYG.

2.2.1. Methodology

In France, values for the socio-economic valuation of public investments are chosen based on official documents. For mortality, the most recent version (Quinet, 2013) uses the results of a set of international works prepared under the aegis of the OECD (Lindhjem *et al.*, 2011; OCDE, 2012). These documents are based on a meta-analysis of 856 valuations of the VPF worldwide in reference to 76 stated preference studies, and the VPF that were proposed

2. The most standard terminology is Value of a Statistical Life (VSL). However, we prefer VPF, in accordance with Desaignes *et al.* (2011) who explain in their first footnote that “the traditional term “value of statistical life” (VSL) is unfortunate, because it tends to evoke hostile reactions by non-economists. However, people tend to accept the concept if it is presented as the “willingness-to-pay for avoiding an anonymous premature death”, i.e. the value of preventing a fatality (VPF)”. Recently, “value of reduced mortality risk” has also been suggested (Simon *et al.*, 2019).

3. The use of an indicator taking into account the quality of years of life (QALY, quality adjusted life years) is not considered here for two reasons. On the one hand, there is no knowledge regarding the quality of life at the time of death and this would require strong hypotheses in order to be established. On the other hand, we consider that these indicators are still not used very much for environmental valuation and are not subject to an international scientific consensus (Cerema, 2016).

Box – Reminders of the Methods for Economic Valuation of Mortality

The economic valuation of mortality is based on three main methods:

- The market price method – often inappropriately called the human capital method – assumes that the value associated with the life of an individual is equal to the future production losses occasioned by their death, with such losses being measured by the value of future revenue discounted based on life expectancy at the age of death. Although easy to implement, it is barely used any more as it does not take into account individual preferences; the value of an individual is represented solely by their production measured by revenue from labour and is very sensitive to the choice of discount rate.
- The revealed preference method is based on situations in which individuals reveal their preferences when choosing consumer goods, implying a trade-off between a market good and a death risk variation. It relies on markets where the death risk level represents one of the characteristics behind the decision: labour markets, housing markets or protection expenditure. The advantage of this method is its reliance on real, observed choices resulting from individual decisions. Disadvantages include the difficulty in isolating the drop in a particular risk when different risks are reduced simultaneously (injury, property loss, drawbacks of a specific job) and the assumption of complete and perfect knowledge of goods, associated risks, the effect of risk attributes on the probability of death etc. In addition, the sample used may not be representative of the general population, under- or over-representing certain groups (workers, owners, etc.). This method is still used to assess the Value of Prevented Fatality (VPF), in particular by the various US federal agencies.
- The stated preference method uses surveys conducted on a sample of the population, which elicit Willingness To Pay (WTP) in order to reduce the probability of death on the basis of hypothetical scenarios. A VPF or Value of a Life Year (VOLY) is then calculated directly. This method is easy to deploy, offers a very accurate description of the trade-off between WTP and the health risk involved, and requires a simpler theoretical framework than that needed for the revealed preference method. The main pitfalls are the various sources of bias and errors that may not always be controlled (see Mitchell & Carson (1989) for an exhaustive presentation and McFadden & Train (2017) for a more critical approach). This method is increasingly used in mortality valuation, particularly by the European agencies.

(taking into account the level of wealth in each country) were largely taken up by the national and supranational bodies in charge of health-environmental valuation. Abroad, the World Bank (World Bank, 2020), the European Union (European Commission, 2020), the WHO and the OECD (WHO-OECD, 2015) have used them in evaluating the health effects of atmospheric pollution.

Quinet (2013) therefore puts forward a single VPF of €3 million₂₀₁₀ for France, considered as a reference, used in the French legislative and regulatory context of the normative framework for the economic valuation of major transport infrastructure projects. He also derives a single VOLY of €115,000₂₀₁₀, on the basis of an average age of the French population of 40 years, and an annual discount rate of 2.5%. This value, like the VPF, depends neither on the scope of application nor on the cause of death.

However, an important finding in the studies based on stated or revealed preferences is that the VPF depends on the context in which death occurs – the nature and level of the underlying risk, age, quality of life and the state of health at death (Chestnut & De Civita, 2009; OECD, 2012; Rabl *et al.*, 2014; Narain & Sall, 2016) – and even the scenario used (Ami *et al.*, 2013). The context of the underlying mortality risk is thus a pertinent factor explaining the extent of the VPF (Hammit, 2007). Ideally, valuations of the VPF and VOLY should be specific to the context of atmospheric pollution.

2.2.2. Choice of Economic Valuation Parameters

- Direct Estimate of a Contextual VOLY:

A review of the literature finds six European stated preference studies where the scenario explicitly mentions exposure to atmospheric pollution as being the origin of the risk of death. Chronologically, Soguel & van Griethuysen (2000) use a sample of Swiss respondents to estimate an implicit VOLY based on a scenario eliciting the WTP for a gain of one hour of life per year. Their estimate of 53,000 Swiss francs (€29,000₂₀₀₈) is calculated as 24×365 times the value of an hour of life. In a scenario based on health risks associated with atmospheric pollution, Chilton *et al.* (2004) estimate the average VOLY for a normal state of health at €45,000 (£27,600) for a sample of UK residents. For a sample of Swiss citizens, Jeanrenaud & Marti (2007) obtain an average VOLY of between €31,000 and €58,000 depending on the scenarios. Desaignes *et al.* (2011) take an

approach similar to that of Chilton *et al.* (2004), based directly on an increase in life expectancy for nine European countries within the framework of the NEEDS programme. Taking the average values for an increase of three months, they recommend a VOLY of €41,000₂₀₀₆ for the EU15 countries plus Switzerland. In Greece, Vlachokostas *et al.* (2011) estimate a VOLY of €41,000 based on a contingent valuation survey eliciting the WTP for an increase in life expectancy of one year thanks to the deployment of air quality improvement measures. Finally, across a sample of French citizens, Chanel & Luchini (2014) express the reduction of mortality as a gain in life years. Considering the VPF as a flow of VOLY discounted at the annual rate of 6.8% (rate estimated in the model based on responses), they derive an average VOLY of €165,000. This value – which is relatively high – is explained by the high discount rate used.

- Choice of a VOLY:

Depending on how a VOLY is obtained (by direct estimate in a contextual stated preference study or by derivation based on a single VPF), the values vary by around a factor of two. As there is no scientific consensus favouring either approach, and as we do not want to favour either, we have chosen the arithmetic average (rounded) of the VOLY adopted by Desaignes *et al.* (2011) and of that recommended in Quinet (2013), i.e. €85,000₂₀₂₀. We note that this value is consistent with that recommended by the British government (£60,000₂₀₁₀ equating to €79,999₂₀₂₀, cf. HM Treasury, 2020) or the EU (€70,000, cf. European Commission, 2020).

- Choice of Discount Rate:

We are taking as our central value the annual risk-free discount rate of $\delta = 2.5\%$ currently favoured in France (Quinet, 2013). It is comparable to the rate of 3% used by the US EPA (2021) to take account of death flows occurring in the future.⁴

2.3. Accounting for Uncertainties

We adopt two approaches. On the one hand, an independent valuation of the uncertainties in the results tables. We account for epidemiological uncertainties based on central estimates using the 95% confidence intervals proposed

4. We are also estimating the sensitivity of economic impacts to the choice of an annual rate of 7%. This choice is based on US EPA (2021, p. F-8) which advocates, in the absence of arbitrage at federal level, performing an economic valuation of health benefits on the basis of 3% (which it recommends) and 7% (as recommended by the Office of Management and Budget, OMB).

in Medina *et al.* (2021). The uncertainties concerning τ , δ and VOLY will be represented by an interval adjusted to their central values, 1 and 5 respectively for τ , 1.5% and 3.5% for δ , and €85,000₂₀₂₀ \pm 33% (namely €56,666 and €113,333) for the VOLY.

On the other hand, we represent a joint valuation of the uncertainties on a figure. It takes into account all of the sources in an integrated approach, using Monte-Carlo simulations (Burmester & Anderson, 1994; CAFE, 2005; Ostro *et al.*, 2006). The epidemiological uncertainty with regard to the exposure-response ratio is accounted for thanks to random draws in a normal distribution whose mean is the central estimate of LYG and whose standard deviation is derived from the 95% CI. For the other uncertainties, we use a triangular distribution defined from the central values and the lower and upper values referenced above for each parameter (τ , δ and VOLY). We then generate 10,000 independent Monte-Carlo replications from these probability distributions, each constituting a monetary valuation. A probability distribution of the economic valuation of the impact of activity restrictions on mortality is then obtained for each of the two pollution indicators (PM_{2.5} and NO₂).

3. Results

The results are set out below for both indicators and must not be added together in order to avoid double counting, as some underlying health effects are common.

3.1. Evaluation of Health Effects

Table 1 presents the results in terms of LYG for various values of τ , by gender and by pollution indicator. The value “close to 0” allows

these results to be compared to those based on the HIA (Medina *et al.*, 2021), which reflect the difference between the health effects (considered immediate) due to ambient air pollution with and without lockdown measures.

For a τ which is close to 0 and for PM_{2.5}, the total numbers of LYG are comparable between the approach including latency (26,313) and the standard HIA approach (27,815). However, the difference is greater for NO₂: 7205 vs. 11,263 for the HIA. This is explained by the fact that the distribution of the population by level of exposure is much more finely measured in the HIA (it is carried out at commune level) than in our approach (based on a weighted national average). It thus allows better consideration of urban exposure, mainly linked to motor traffic (principal source of NO₂) and affecting a large proportion of the population (60% of the population live in an urban unit of more than 20,000 inhabitants, Medina *et al.*, 2021, Table 3).

When τ increases, the total number of LYG drops for both pollution indicators, for two reasons. The main reason stems from the decrease in impacts seen in the first year, linked to the lower RR attained as a result of lockdown (see Figure II), a phenomenon which develops as the cohort ages. This is illustrated in Figure III, which represents the distribution over time of LYG following lockdown, for the three values of τ used in our analysis (for PM_{2.5}). The second, more ancillary reason, is explained by the other reasons for death (independent of exposure to air pollution) which affect the ageing of the cohort. Their contribution to its extinction becomes more significant as the evolution of RR towards RR_{NE} following lockdown, is slow (high τ), reducing the total number of LYG attributed to

Table 1 – Total number of life years gained long-term following lockdown

Values of τ	PM _{2.5}			NO ₂		
	Men	Women	Total	Men	Women	Total
Close to 0	14,425 (5,266–22,118)	11,888 (4,340–18,228)	26,313 (9,606–40,346)	3,950 (1,394–6,269)	3,255 (1,149–5,166)	7,205 (2,543–11,435)
1	9,118 (3,329–13,982)	7,515 (2,743–11,523)	16,633 (6,072–25,505)	2,497 (881–3,963)	2,058 (726–3,266)	4,555 (1,607–7,229)
3	4,089 (1,493–6,270)	3,370 (1,230–5,167)	7,459 (2,723–11,437)	1,120 (395–1,777)	923 (326–1,464)	2,043 (721–3,241)
5	2,615 (955–4,009)	2,155 (787–3,304)	4,770 (1,742–7,313)	716 (253–1,136)	590 (208–936)	1,306 (461–2,072)
7	1,920 (701–2,944)	1,583 (578–2,427)	3,503 (1,279–5,371)	526 (186–835)	433 (153–688)	959 (339–1,523)
HIA (2021)			27,815 (9,709–44,414)			11,263 (3,946–17,995)

Notes: The figures in brackets are established based on the 95% CI of the health data.

Sources: Calculation by the author and Medina *et al.* (2021).

the drop in exposure. This contribution is only marginally offset by a slower return of RR to the level of RR_E when τ increases (cf. Figure II). These two reasons thus explain why the discrepancies between our results and those of the HIA widen as τ increases, irrespective of the pollution indicator (cf. Table 1). For the central value $\tau=3$, they are thus lower by a factor of 3.7 (for $PM_{2.5}$) and 5.5 (for NO_2).

3.2. Economic Results

3.2.1. Independent Processing of Uncertainties

Table 2 presents the discounted monetary valuation of the flow of LYG for the three values of τ , δ and VOLY used to reflect uncertainty. For the central values of these parameters it is €504 million (184-773) for $PM_{2.5}$, and

Figure III – Distribution over time of the number of life years gained (LYG) following lockdown, function of τ (For $PM_{2.5}$)

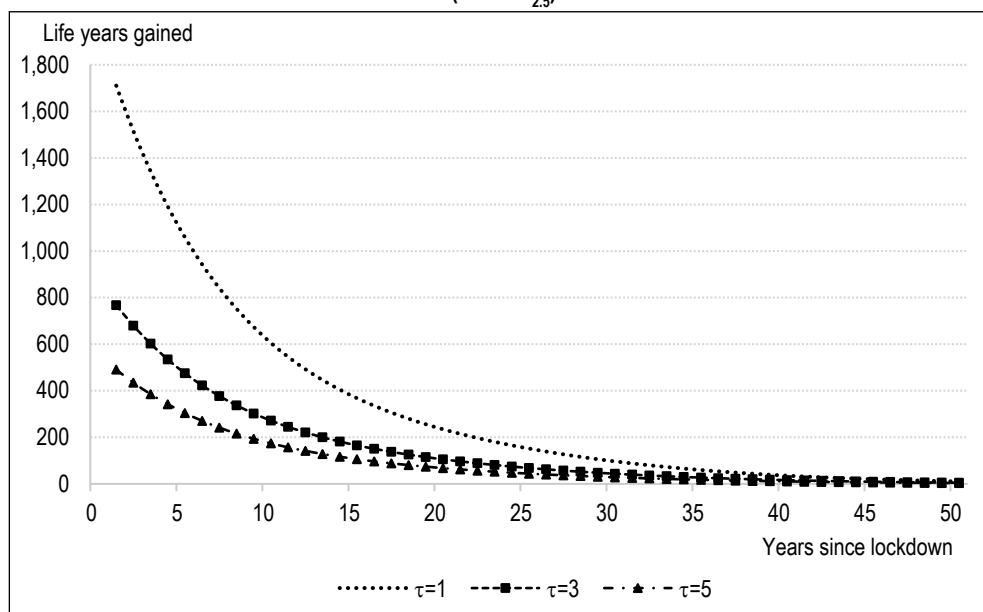


Table 2 – Discounted monetary valuation of the total number of life years gained long-term following lockdown (in € Millions)

		$PM_{2.5}$			NO_2		
τ	δ	VOLY			VOLY		
		€56,667	€85,000	€113,333	€56,667	€85,000	€113,333
1	1.5%	816 (298–1,251)	1,224 (447–1,877)	1,632 (596–2,503)	223 (79–355)	335 (118–532)	447 (157–709)
	2.5%	749 (273–1,149)	1,124 (410–1,723)	1,499 (547–2,297)	205 (73–325)	308 (109–488)	411 (145–651)
	3.5%	693 (253–1,062)	1,039 (379–1,593)	1,385 (505–2,124)	189 (67–301)	284 (100–451)	379 (133–601)
3	1.5%	366 (133–561)	549 (200–842)	732 (267–1,123)	100 (35–159)	150 (53–239)	200 (71–319)
	2.5%	336 (123–515)	504 (184–773)	672 (245–1,031)	92 (33–146)	138 (49–219)	184 (65–292)
	3.5%	311 (113–476)	466 (170–714)	621 (227–952)	85 (30–135)	128 (45–202)	171 (60–269)
5	1.5%	234 (85–359)	351 (128–538)	468 (171–717)	64 (32–102)	96 (34–153)	128 (45–204)
	2.5%	215 (79–329)	322 (118–494)	429 (157–659)	59 (21–93)	88 (31–140)	117 (41–187)
	3.5%	199 (73–305)	298 (109–457)	397 (145–609)	55 (19–86)	82 (29–129)	109 (39–172)
HIA (2021)		1,576 (550–2,517)	2,364 (825–3,775)	3,152 (1,100–5,033)	638 (223–1,020)	957 (335–1,530)	1,276 (447–2,040)

Notes: The figures in brackets are established based on the 95% CI of the health data.
Sources: Author's calculations.

€138 million (49-219) for NO₂. When we compare them to the monetary valuations calculated based on the results of Medina *et al.* (2021) and presented in the last line, they are 4.7 (for PM_{2.5}) and 6.9 (for NO₂) times lower, reflecting the combined effect of latency and discounting. For a given value of τ or VOLY, the results are not particularly sensitive to the value of the discount rate, which is explained by the fact that the flow of LYG, decreasing over time, limits the impact of discounting.⁵ The results are proportional to the VOLY, *ceteris paribus*. On the other hand, the choice of τ is more determining: the move from a value of 1 to 5 divided the monetary valuation by 4 approximately, for both pollution indicators.

3.2.2. Joint Processing of Uncertainties

Figure IV represents the distribution of monetary valuations jointly considering the different sources of uncertainties, based on 10,000 Monte-Carlo replications. It gives rise to an average value and empiric 95% CI of €708 million (151-1,678) for PM_{2.5} and 193 million (38-462) for NO₂, approximately 40% more than the central values of Table 2. This difference is mainly explained by the non-linear impact of τ on the valuation, favouring higher values due to the random draws from a triangular distribution. The difference actually sits at less than 8% when calculated based on the averages of the 27 central values ($3\delta \times 3\tau \times 3VOLY$) of Table 2, i.e. €653 million (PM_{2.5}) and €179 million (NO₂). It thus approaches those obtained in other studies

comparing independent vs. joint processing of uncertainties (Adélaïde *et al.*, 2021a; Chanel *et al.*, 2014).

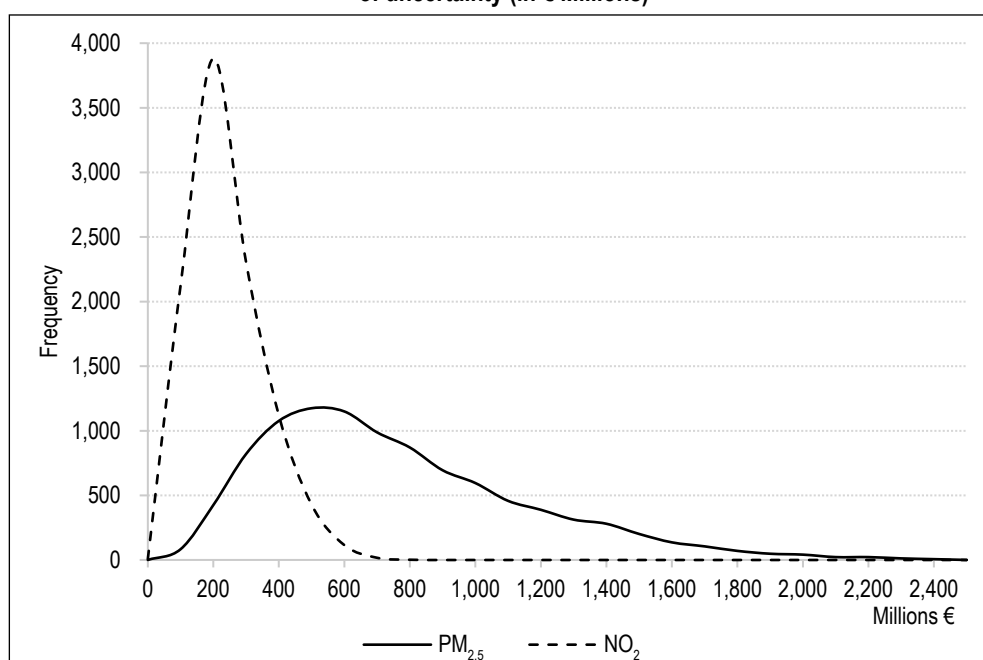
* *
*

In terms of public health, our results confirm the importance of reducing – even temporarily and by a low amount – the exposure of the population to atmospheric pollution. The standard approach based on difference evaluates the effects associated with long-term mortality at €2.4 billion for PM_{2.5} and €957 million for NO₂. Taking into account latency (and discounting future LYG flows), our recommended approach involves dividing these values by 5 approximately for PM_{2.5} (i.e. €500 million) and 7 for NO₂ (i.e. €140 million). Thus it is crucial to be aware of the implicit epidemiological choices associated with these approaches when they are included in the economic analysis.

It is difficult to make a direct comparison between the monetary valuations that we have obtained and those from literature, for two reasons. Firstly, the works on activity restrictions caused by COVID-19 have only just begun to be circulated and published. Secondly, the

5. Thus the valuations performed using the annual discount rate of 7% (recommended by the US OMB) represent approximately 73% of the values obtained with the rate of 2.5%, *ceteris paribus*.

Figure IV – Distribution of the total discounted monetary valuation jointly considering the different sources of uncertainty (in € Millions)



valuations depend on the methodology used (modelling and comparison of levels, regression approaches or application of RR, period of restriction studied), exposure measured (choice of pollution indicators, calculation of exposure values), epidemiological choices (RR, reference scenario), measurement of mortality gains (premature deaths avoided or LYG), and the choice of monetary values.

However, some studies have evaluated the impact on long-term mortality of the drop in atmospheric pollution linked to activity restrictions, and offer comparative data.

Assuming an immediate resumption of activity for the whole of 2020 following lockdown, Giani *et al.* (2020) estimate that 76,400 (62,600-86,900) premature deaths would have been avoided in China and 13,600 (11,900-15,300) in Europe, including around 1,250 in France (see Figure S5 of their appendix). Assuming a lockdown throughout 2020, Hao *et al.* (2021) estimate the drop in average concentration of PM_{2.5} to be 32.2% for China (compared with 2015-2019) with the number of deaths avoided being 140,200 (122,200-156,000). By way of perspective, we note that Medina *et al.* (2021) evaluate the drop in long-term mortality in France linked to the total elimination of the anthropic portion of atmospheric pollution at, respectively, 491,800 LYG (171,900-784,800) per year for PM_{2.5} and 106,400 (37,300-169,900) for NO₂. In economic terms, this represents respectively €42 billion and €9 billion per year.

Some limits need to be specified. First of all, the transposition of a negative exponential function obtained from smoking cessation to a reduction in exposure to atmospheric pollution most likely depends – in addition to the similar exposure routes and target organs – on the nature of the chemicals involved, biokinetics, bioaccumulation, and the extent and temporality of the reduction. We note, however, that the negative exponential function of equation (1) is also adapted to reflect the phenomena of degradation in disciplines other than health (such as physics, biology, etc.), that it is compatible with the

literature analysis carried out by Walton (2010), and that the broad interval used for τ reflects the uncertainty linked to this transposability.

The analysis could then be refined. On the one hand, we use a dynamic cohort based on an average variation in the exposure of the population over time. The use of exposure variations modelled at local level (grid measuring 4 km by 4 km) and their overlay with the communal population data should allow local specificities to be better taken into account, and mortality tables covering a more disaggregated level than national to be used. On the other hand, part of the population has moved out of urban areas into more rural environments (approximately 1.4 million, including 450,000 from Paris, according to Galiana *et al.*, 2020). As the exposure levels in more urbanised areas are higher than those in rural environments, in particular for NO₂ (Medina *et al.*, 2021), the effect of lockdown on mortality is undoubtedly underestimated in the population. Remote working has also contributed to reducing the exposure of the population concerned.

Lastly, mortality is evaluated monetarily on the basis of preferences stated by the population and not on an observation of market prices. These preferences represent the expression of a willingness to pay to reduce the probability of death, and include non-market components. The valuation of mortality also represents losses of collective well-being, therefore essentially a non-market component, for which a direct comparison with purely market components (such as the gross domestic product) is not recommended.

Finally, we note that, in addition to the drop in mortality following the impact of activity restrictions on the concentrations of PM_{2.5} and NO₂, there are gains in morbidity linked to the respiratory or cardiovascular impacts (see Venter *et al.*, 2021, for paediatric asthma for example). However, potential negative health effects are also associated, since some studies demonstrate an increase in ozone levels and the associated mortality (Liu *et al.*, 2021; Venter *et al.*, 2021). □

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